

TESTING FOR A DISEASE

- Suppose 1% of the population has a disease
- there is a diagnostic test, that finds it 80% of the time
(assuming the subject has it)
- the test also produces false positives, at a rate of 9.6%
(you're fine, but the test says you're not)

If someone tests positive,
what are the odds that they have the disease?

4 events

	1% Have disease	99% Don't have disease
Test :	80%	9.6%
Test :-	20%	90.4%

$$P(\text{disease} | \text{test :}) = \frac{P(\text{disease} \cap \text{test :})}{P(\text{test :})} = \frac{P(\text{test :} | \text{disease}) \cdot P(\text{disease})}{P(\text{test :})}$$

$$P(\text{test :}) = \begin{cases} P(\text{test :} | \text{disease}) \cdot P(\text{disease}) \\ + \\ P(\text{test :} | \text{no disease}) \cdot P(\text{no disease}) \end{cases} = 0.8 \cdot 0.01 = 0.008 \text{ (0.8\%)} \\ P(\text{test :} | \text{no disease}) \cdot P(\text{no disease}) = 0.096 \cdot 0.99 \approx 0.095 \text{ (9.5\%)}$$

$$P(\text{disease} | \text{test :}) = \frac{0.008}{0.008 + 0.095} \approx 7.8\%$$

Bayes theorem

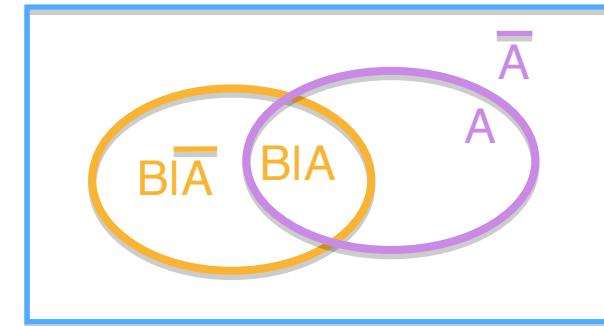
$$P(A \cap B) = P(A|B) \cdot P(B)$$

$$= P(B|A) \cdot P(A)$$



$$P(A|B) = P(B|A) \cdot \frac{P(A)}{P(B)}$$

$$P(A|B) = \frac{P(B|A) \cdot P(A)}{P(B|A) \cdot P(A) + P(B|\bar{A}) \cdot P(\bar{A})}$$



$$P(A|B) = \frac{P(B|A) \cdot P(A)}{P(B|A) \cdot P(A) + P(B|\bar{A}) \cdot P(\bar{A})}$$

A: have disease
 B: test \approx

	1% A	99% \bar{A}
$B A = 80\%$		$B \bar{A} = 9.6\%$

$$\begin{aligned} P(A) &= 0.01 \\ P(\bar{A}) &= 0.99 \end{aligned}$$

$$\begin{aligned} P(B|A) &= 0.8 \\ P(B|\bar{A}) &= 0.096 \end{aligned}$$

$$P(A|B) = \frac{0.8 \cdot 0.01}{0.8 \cdot 0.01 + 0.096 \cdot 0.99}$$

RANDOM VARIABLES

A quote from your textbook:

"A random variable is neither random nor variable"

We have been using random variables, implicitly.

ex: roll 2 dice, examine probability that sum = k, or = even.

↳ define random variable X : sum of two dice rolls.

$$\text{So, } X[(1,2)] = 3$$

$$X[(5,5)] = 10$$

↳ define random variable Y : parity of two dice rolls.

$$\text{So, } Y[(1,2)] = 1$$

$$Y[(5,5)] = 0$$

or $= 13$
 $= -22$ } (arbitrary)

Think of a r.v. as a function,
mapping sample space to whatever you like
usually a number

Then we can express questions neatly:

$$P(X < 3) = \frac{1}{36} \quad \begin{matrix} 2 \text{ dice} \\ \leftarrow \text{sum} \end{matrix}$$
$$P(Y=1) = \frac{1}{2} \quad \begin{matrix} \\ \leftarrow \text{parity} \end{matrix}$$

We can also eliminate absurd events, e.g., $P(X=13)=0$

Section 33 discusses in more detail: not required for our class

EXPECTATION : the very basics

As mentioned, a r.v. X can have several values.

It is based on outcomes that result from a random process.

So we don't know what value it will have.

But we can expect it to have some value

Expected value = weighted average

(e.g. average parity for unfair coin = $\frac{1}{2}$; weighted average $\neq \frac{1}{2}$)

Expected value = weighted average

$$E(X) = \sum_{*} y \cdot P(X=y)$$

*

} see example 34.3 in book (mainly p.237)

↳ * over all possible values y , compatible with X .

(however we only care about finitely many)

$$E(X) = \sum_{s \in S} [X(s) \cdot P(s)]$$

}

e.g. roll 1 die. $X = \text{number observed}$.

$$E(X) = \sum_{i=1}^6 X(i) \cdot P(i)$$

$$= 1 \cdot \frac{1}{6} + 2 \cdot \frac{1}{6} + 3 \cdot \frac{1}{6} + 4 \cdot \frac{1}{6} + 5 \cdot \frac{1}{6} + 6 \cdot \frac{1}{6}$$

$$= \frac{1}{6} \cdot (1+2+3+4+5+6) = 3.5$$

over all samples
that define X . } a finite number

$$E(X) = \sum y \cdot P(X=y)$$

| typically use this when you have a good way of aggregating outcomes

example: roll 2 dice. $X = |\text{difference between the 2}|$

possible values of $X \rightarrow 0 \quad 1 \quad 2 \quad 3 \quad 4 \quad 5$

outcomes supporting value $\rightarrow 6 \quad 5 \cdot 2 \quad 4 \cdot 2 \quad 3 \cdot 2 \quad 2 \cdot 2 \quad 1 \cdot 2$
(for probability, divide by 36)

$$E(X) = \frac{0 + 10 + 16 + 18 + 16 + 10}{36}$$

(~ 1.944)

EXPECTATION : PROPERTIES

LINEARITY OF EXPECTATION (important)

↪ $c_1, c_2 \in \mathbb{R}$ $E(c_1 X + c_2 Y) = c_1 \cdot E(X) + c_2 \cdot E(Y)$

Generally, $E(c_1 X_1 + c_2 X_2 + \dots + c_n X_n) = c_1 E(X_1) + c_2 E(X_2) + \dots + c_n E(X_n)$

$E(\sum c_i X_i) = \sum c_i E(X_i)$ → Does NOT assume independence

Independence: $P(X=a \& Y=b) = P(X=a) \cdot P(Y=b)$
for all $a, b \dots$ (def. 33.6)

2 dice, A, B. $X = \text{result of A.}$ $Y = \text{result of B.}$ $Z = X+Y$

$$E(Z) = E(X+Y) = E(X) + E(Y) = 2 \cdot 3.5 = 7$$

EXPECTATION : PROPERTIES

$$E(X+Y) = E(X) + E(Y)$$

Linearity of expectation doesn't assume independence

but $E(X \cdot Y) \neq E(X) \cdot E(Y)$ in general.

If X & Y are independent, then $E(X \cdot Y) = E(X) \cdot E(Y)$

However, $E(X \cdot Y) = E(X) \cdot E(Y)$ does NOT imply
 X & Y are independent.

(see example 34.15)

- We are skipping the proofs of most statements in this section.
- You are not required to study these, but it would probably be beneficial.
- We are also skipping variance, which is an important concept to learn independently.

INDICATOR RANDOM VARIABLES

(taking value 0 or 1)

We already saw this : Y : parity of rolling one die.

Another example: flip a coin 10 times.

$X = \#\text{times we see pattern HT}$

$$E(X) = ?$$

Try solving this "directly"

INDICATOR RANDOM VARIABLES

flip a coin 10 times. $X = \#\text{times we see pattern HT}$

HT could appear at flips 1&2, or 2&3, ..., or 9 & 10

Define r.v. $X_i = \begin{cases} 1 & \text{if flips } i \text{ & } i+1 \text{ produce HT} \\ 0 & \text{otherwise} \end{cases}$

$$X = X_1 + X_2 + \cdots + X_9$$

Notice X_1 & X_2 are
not independent. $P(X_i=1) = \frac{1}{4}$
 $P(X_1 \wedge X_2) = 0$

$$E(X) = E(X_1 + X_2 + \cdots + X_9)$$

$$= E(X_1) + E(X_2) + \cdots + E(X_9)$$

linearity of expectation

$$E(X_i) = 0 \cdot P(X_i=0) + 1 \cdot P(X_i=1) = \frac{1}{4}$$

$$= 9 \cdot \frac{1}{4}$$

(p.240)